# Practical Numerical <br> <br> Training UKNum <br> <br> Training UKNum Statistics, data modeling 

Dr. C. Mordasini

Max Planck Institute for Astronomy, Heidelberg

Program:

1) Elementary statistics
2) Regression analysis
3) Linear regression
4) Non-linear regression

1 Elementary statistics

## Basic statistical quantities I

- Study a statistical sample containing n values e.g. measurements. -We can characterise the sample by so called moments, i.e. sums of integer powers of the values.


## Mean

Estimates the value around which central clustering occurs

$$
\bar{y}=\frac{\sum_{i=1}^{n} y_{i}}{n}
$$

Alternatives are the median or the mode.

## Range

$$
R=y_{\max }-y_{\min }
$$

$y_{\text {max }}$ is the maximum of the values of $y_{i}, i=1,2, \ldots, n$, $y_{\text {min }}$ is the minimum of the values of $y_{i}, i=1,2, \ldots, n$.

Problem: outliers

## Basic statistical quantities II

## Residual

The residual between a data point i and the mean is the residual i

$$
e_{i}=y_{i}-\bar{y}
$$

The residual can be negative or positive and hence if one calculates the sum of such differences to find the overall spread, the differences may simply cancel each other. That is why the sum of the square of the differences is a better measure.

## Summed squared error

$$
S_{t}=\sum_{i=1}^{n}\left(y_{i}-\bar{y}\right)^{2}
$$

The magnitude of the summed squared error is dependent on the number of data points. Therefore, we want an average.

## Basic statistical quantities III

## Variance

An average value of the summed squared error is the variance

$$
\sigma^{2}=\frac{S_{t}}{n-1}=\frac{\sum_{i=1}^{n}\left(y_{i}-\bar{y}\right)^{2}}{n-1}
$$

The variance is divided by $(\mathrm{n}-1)$ and not n because with the use of the mean in calculating the variance, we lose the independence of one of the data points. That is, if you know the mean of $n$ data points, then the value of one of the $n$ data points can be calculated by knowing the other ( $\mathrm{n}-1$ ) values. N should be used if the mean is known externally (not calculated from the sample).

## Basic statistical quantities IV

## Variance continued

There are different ways of writing the variance numerically like

$$
\operatorname{Var}\left(x_{1} \ldots x_{N}\right)=\frac{1}{N-1}\left[\left(\sum_{j=1}^{N} x_{j}^{2}\right)-N \bar{x}^{2}\right] \approx \overline{x^{2}}-\bar{x}^{2}
$$

One which reduces round-off errors (large samples) is the corrected two pass algorithm. First calculate the mean, then the variance as

$$
\operatorname{Var}\left(x_{1} \ldots x_{N}\right)=\frac{1}{N-1}\left\{\sum_{j=1}^{N}\left(x_{j}-\bar{x}\right)^{2}-\frac{1}{N}\left[\sum_{j=1}^{N}\left(x_{j}-\bar{x}\right)\right]^{2}\right\}
$$

## Basic statistical quantities V

## Standard deviation

To bring the variation back to the same level of units as the original data, the standard deviation is defined as

$$
\sigma=\sqrt{\frac{S_{t}}{n-1}}=\sqrt{\frac{\sum_{i=1}^{n}\left(y_{i}-\bar{y}\right)^{2}}{n-1}}
$$

## Coefficient variation

The ratio of the standard deviation to the mean, known as the coefficient variation is used to normalize the spread of a sample

$$
c . v=\frac{\sigma}{\bar{y}} \times 100
$$

## Basic statistical quantities VI

## Skewness

Also known as the third moment, skewness characterizes the degree of asymmetry of a distribution around its mean. It is a non-dimensional number.

$$
\operatorname{Skew}\left(x_{1} \ldots x_{N}\right)=\frac{1}{N} \sum_{j=1}^{N}\left[\frac{x_{j}-\bar{x}}{\sigma}\right]^{3}
$$



A positive value signifies a distribution
 with an asymmetric tail extending out towards more positive $x$; a negative value signifies a distribution whose tail extends out towards more negative x .

## Basic statistical quantities VII

## Kurtosis

Also known as the forth moment, the kurtosis characterizes the relative peakedness of flatness of a distribution. It is a non-dimensional number.

$$
\operatorname{Kurt}\left(x_{1} \ldots x_{N}\right)=\left\{\frac{1}{N} \sum_{j=1}^{N}\left[\frac{x_{j}-\bar{x}}{\sigma}\right]^{4}\right\}-3
$$



A distribution with positive kurtosis is termed leptokurtic. The Matterhorn is an example. A distribution with negative kurtosis is termed platykurtic. An inbetween distribution is mesokurtic. The reference is a normal distribution.

High moments (skew, kurt) are less robust than the mean or std. deviation.

## Basic statistical quantities VIII

## Median

The median of a probability distribution function $p(x)$ is the value $x_{\text {med }}$ for which larger and smaller values of x are equally probable:

$$
\int_{-\infty}^{x_{\mathrm{med}}} p(x) d x=\frac{1}{2}=\int_{x_{\mathrm{med}}}^{\infty} p(x) d x
$$

The median of a sample of values $x_{1}, \ldots, x_{N}$ is that value $x_{i}$ which has equal numbers of values above it and below it. Of course, this is not possible when $N$ is even. In that case it is conventional to estimate the median as the mean of the unique two central values. If the values are sorted into ascending order, then the formula for the median is

$$
x_{\mathrm{med}}= \begin{cases}x_{(N+1) / 2}, & N \text { odd } \\ \frac{1}{2}\left(x_{N / 2}+x_{(N / 2)+1}\right), & N \text { even }\end{cases}
$$

## Basic statistical quantities IX

## Mode

The mode of a probability distribution function $p(x)$ is the value of $x$ where $p$ takes on a maximum value. The mode is useful primarily when there is a single, sharp maximum, in which case it estimates the central value.

Occasionally, a distribution will be bimodal, with two relative maxima; then one may wish to know the two modes individually. Note that, in such cases, the mean and median are not very useful, since they will give only a "compromise" value between the two peaks.

2 Regression analysis

## Regression analysis

What is regression analysis?
Regression analysis gives quantitative information on the relationship between a response (dependent) variable and one or more predictor (independent) variables to the extent that the necessary information is contained in the data.

What is regression analysis used for?

1. prediction
2. model adaption and
3. parameter estimation.

In other words: fitting

## Least squares method

This is the most popular method of parameter estimation for coefficients of regression models. It has well known probability distributions and gives unbiased estimators of regression parameters with the smallest variance.

We wish to predict the response to $n$ data points $\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots \ldots$, $\left(\mathrm{x}_{\mathrm{n}}, \mathrm{y}_{\mathrm{n}}\right)$ by a regression model given by

$$
y=f(x)
$$

where the function $f(x)$ has regression constants that need to be estimated. For example
$f(x)=a_{0}+a_{1} x$ is a straight-line regression model with constants $a_{0}$ and $a_{1}$
$f(x)=a_{0} e_{1} x$ is an exponential model with constants $a_{0}$ and $a_{1}$
$f(x)=a_{0}+a_{1} x+a_{2} x^{2}$ is a quadratic model with constants $a_{0}, a_{1}$ and $a_{2}$

## Least squares method II

A measure of goodness of fit, that is how the regression model $f(x)$ predicts the response variable y is the magnitude of the residual, $\mathrm{E}_{\mathrm{i}}$ at each of the n data points.

$$
E_{i}=y_{i}-f\left(x_{i}\right), i=1,2, \ldots n
$$

Ideally, if all the residuals $\mathrm{E}_{\mathrm{i}}$ are zero, one may have found an equation in which all the points lie on a model.

In the least squares method, estimates of the constants of the models are chosen such that minimization of the sum of the squared residuals is achieved, that is minimize

$$
\dot{N}_{\tilde{m}^{2}}^{E^{2}}
$$

from where the name least squares.

## Least squares method III

In other words, given n data points $\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots,\left(x_{n}, y_{n}\right)$ best fit $y=f(x)$ to the data. The best fit is based on minimizing the sum of the square of the residuals, $S_{r}$.

Residual at a point is

$$
\varepsilon_{i}=y_{i}-f\left(x_{i}\right)
$$

Sum of the square of the residuals

$$
S_{r}=\sum_{i=1}^{n}\left(y_{i}-f\left(x_{i}\right)\right)^{2}
$$



# 3 Linear Regression 

## Linear Regression

Given n data points $\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots,\left(x_{n}, y_{n}\right)$ best fit

$$
y=a_{0}+a_{1} x
$$

to the data.


Illustration with Mathematica

## Parameter determination I

The least squares criterion minimizes the sum of the square of the residuals in the model, and produces a unique line.

$$
S_{r}=\sum_{i=1}^{n} \varepsilon_{i}^{2}=\sum_{i=1}^{n}\left(y_{i}-a_{0}-a_{1} x_{i}\right)^{2}
$$

Our task is to minimize the sum of the square of the residuals by determining the best regression (or fit) parameters $a_{0}$ and $a_{1}$.

At a minimum, we must have (using the chain rule):

$$
\begin{aligned}
& \frac{\partial S_{r}}{\partial a_{0}}=-2 \sum_{i=1}^{n}\left(y_{i}-a_{0}-a_{1} x_{i}\right)(-1)=0 \\
& \frac{\partial S_{r}}{\partial a_{1}}=-2 \sum_{i=1}^{n}\left(y_{i}-a_{0}-a_{1} x_{i}\right)\left(-x_{i}\right)=0
\end{aligned}
$$

## Parameter determination II

This gives

$$
\begin{aligned}
& \sum_{i=1}^{n} a_{0}+\sum_{i=1}^{n} a_{1} x_{i}=\sum_{i=1}^{n} y_{i} \\
& \sum_{i=1}^{n} a_{0} x_{i}+\sum_{i=1}^{n} a_{1} x_{i}^{2}=\sum_{i=1}^{n} y_{i} x_{i}
\end{aligned}
$$

Noting that $\sum_{i=1}^{n} a_{0}=a_{0}+a_{0}+\ldots+a_{0}=n a_{0}$

$$
\begin{aligned}
& n a_{0}+a_{1} \sum_{i=1}^{n} x_{i}=\sum_{i=1}^{n} y_{i} \\
& a_{0} \sum_{i=1}^{n} x_{i}+a_{1} \sum_{i=1}^{n} x_{i}^{2}=\sum_{i=1}^{n} x_{i} y_{i}
\end{aligned}
$$

## Parameter determination III

$$
\begin{aligned}
& n a_{0}+a_{1} \sum_{i=1}^{n} x_{i}=\sum_{i=1}^{n} y_{i} \\
& a_{0} \sum_{i=1}^{n} x_{i}+a_{1} \sum_{i=1}^{n} x_{i}^{2}=\sum_{i=1}^{n} x_{i} y_{i}
\end{aligned}
$$

This later can be understood as a system of two linear equations, which we can write in a $2 \times 2$ matrix form as we have seen in the last lecture, with the unknowns $\mathrm{a}_{0}$ and $\mathrm{a}_{1}$ (all $\mathrm{x}_{\mathrm{i}}$ and $\mathrm{y}_{\mathrm{i}}$ are known). For such a small matrix, one quickly find the result

$$
a_{0}=\frac{\sum_{i=1}^{n} x_{i}^{2} \sum_{i=1}^{n} y_{i}-\sum_{i=1}^{n} x_{i} \sum_{i=1}^{n} x_{i} y_{i}}{n \sum_{i=1}^{n} x_{i}^{2}-\left(\sum_{i=1}^{n} x_{i}\right)^{2}}
$$

$$
a_{1}=\frac{n \sum_{i=1}^{n} x_{i} y_{i}-\sum_{i=1}^{n} x_{i} \sum_{i=1}^{n} y_{i}}{n \sum_{i=1}^{n} x_{i}^{2}-\left(\sum_{i=1}^{n} x_{i}\right)^{2}}
$$

## Parameter determination IV

Defining

$$
\begin{array}{ll}
\bar{x}=\frac{\sum_{i=1}^{n} x_{i}}{n} & \bar{y}=\frac{\sum_{i=1}^{n} y_{i}}{n} \\
S_{x y}=\sum_{i=1}^{n} x_{i} y_{i}-n \bar{x} \bar{y} & S_{x x}=\sum_{i=1}^{n} x_{i}^{2}-n \bar{x}
\end{array}
$$

we can write the least squares linear fit parameters as

$$
a_{1}=\frac{S_{x y}}{S_{x x}} \quad a_{0}=\bar{y}-a_{1} \bar{x}
$$

## Example I

Given the following arbitrary data set, find the best fitting last square model.

| $X$ | Y |
| :---: | :---: |
|  |  |
| 0.698132 | 0.188224 |
| 0.959931 | 0.209138 |
| 1.134464 | 0.230052 |
| 1.570796 | 0.250965 |
| 1.919862 | 0.313707 |



Find the constants $\mathrm{a}_{0}$ and $\mathrm{a}_{1}$ for the linear model given by

$$
y=a_{0}+a_{1} x
$$

## Example II

Using the equations described for $\mathrm{a}_{0}$ and $\mathrm{a}_{1}$ we rapidly find

$$
\begin{aligned}
& \mathrm{a}_{0}=1.1767 \times 10^{-1} \\
& \mathrm{a}_{1}=9.6091 \times 10^{-2}
\end{aligned}
$$



## Nonlinear Regression

Some popular nonlinear regression models:

1. Exponential model:

$$
\begin{aligned}
& \left(y=a e^{b x}\right) \\
& \left(y=a x^{b}\right) \\
& \left(y=\frac{a x}{b+x}\right) \\
& \left(y=a_{0}+a_{1} x+\ldots+a_{m} x^{m}\right)
\end{aligned}
$$

2. Power model:
3. Saturation growth model:

## Nonlinear Regression

Given $n$ data points $\quad\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots,\left(x_{n}, y_{n}\right)$ best fit $y=f(x)$ to the data, where $f(x)$ is a nonlinear function of $x$.


Figure. Nonlinear regression model for discrete y vs. $x$ data

## Exponential Model

Given $\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots,\left(x_{n}, y_{n}\right)$ best fit $\quad y=a e^{b x}$ to the data.


Figure. Exponential model of nonlinear regression for y vs. $x$ data

## Finding Constants of Exponential Model I

The sum of the square of the residuals is defined as

$$
S_{r}=\sum_{i=1}^{n}\left(y_{i}-a e^{b x_{i}}\right)^{2}
$$

Differentiate with respect to $a$ and $b$

$$
\begin{aligned}
& \frac{\partial S_{r}}{\partial a}=\sum_{i=1}^{n} 2\left(y_{i}-a e^{b x_{i}}\right)\left(-e^{b x_{i}}\right)=0 \\
& \frac{\partial S_{r}}{\partial b}=\sum_{i=1}^{n} 2\left(y_{i}-a e^{b x_{i}}\right)\left(-a x_{i} e^{b x_{i}}\right)=0
\end{aligned}
$$

## Finding Constants of Exponential Model II

Rewriting the equations, we obtain

$$
\begin{aligned}
&- \sum_{i=1}^{n} y_{i} e^{b x_{i}}+a \sum_{i=1}^{n} e^{2 b x_{i}}=0 \\
& \sum_{i=1}^{n} y_{i} x_{i} e^{b x_{i}}-a \sum_{i=1}^{n} x_{i} e^{2 b x_{i}}=0
\end{aligned}
$$

## Finding constants of Exponential Model III

Solving the first equation for $a$ yields

$$
a=\frac{\sum_{i=1}^{n} y_{i} e^{b x_{i}}}{\sum_{i=1}^{n} e^{2 b x_{i}}}
$$

Substituting a back into the previous equation

$$
\sum_{i=1}^{n} y_{i} x_{i} e^{b x_{i}}-\frac{\sum_{i=1}^{n} y_{i} e^{b x_{i}}}{\sum_{i=1}^{n} e^{2 b x_{i}}} \sum_{i=1}^{n} x_{i} e^{2 b x_{i}}=0
$$

The constant $b$ can be found through numerical methods such as bisection method. Once it is found, we can also calculate $a$.

## Example - Exponential Model I

Many patients get concerned when a test involves injection of a radioactive material. For example for scanning a gallbladder, a few drops of Technetium-99m isotope is used. Half of the Technetium-99m would be gone in about 6 hours. It however takes about 24 hours for the radiation levels to reach what we are exposed to in day-to-day activities. Below is given the relative intensity of radiation as a function of time.

Relative intensity of radiation as a function of time.

| $\mathrm{t}(\mathrm{hrs})$ | 0 | 1 | 3 | 5 | 7 | 9 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\gamma$ | 1.000 | 0.891 | 0.708 | 0.562 | 0.447 | 0.355 |

## Example - Exponential Model II

The relative intensity is related to time by the equation $\quad \gamma=A e^{\lambda t}$
Find: a) The value of the regression constants $A$ and $\lambda$
b) The half-life of Technetium-99m
c) Radiation intensity after 24 hours


## Constants of the Model

$$
\gamma=A e^{\lambda t}
$$

The value of $\lambda$ is found by solving the nonlinear equation

$$
f(\lambda)=\sum_{i=1}^{n} \gamma_{i} t_{i} e^{\lambda t_{i}}-\frac{\sum_{i=1}^{n} \gamma_{i} e^{\lambda t_{i}}}{\sum_{i=1}^{n} e^{2 \lambda t_{i}}} \sum_{i=1}^{n} t_{i} e^{2 \lambda t_{i}}=0
$$

Once it is found,
A is given as

$$
A=\frac{\sum_{i=1}^{n} \gamma_{i} e^{\lambda t_{i}}}{\sum_{i=1}^{n} e^{2 \lambda t_{i}}}
$$

## Solving the non-linear equation

$$
f(\lambda)=\sum_{i=1}^{n} \gamma_{i} t_{i} e^{\lambda t_{i}}-\frac{\sum_{i=1}^{n} \gamma_{i} e^{\lambda t_{i}}}{\sum_{i=1}^{n} e^{2 \lambda t_{i}}} \sum_{i=1}^{n} t_{i} e^{2 \lambda t_{i}}=0
$$



The value of $A$ can now be calculated also:

$$
A=\frac{\sum_{i=1}^{6} \gamma_{i} e^{\lambda t_{i}}}{\sum_{i=1}^{6} e^{2 \lambda t_{i}}}=0.9998
$$

$$
\lambda=-0.1151
$$

## Plot of data and regression curve

$$
\gamma=0.9998 e^{-0.1151 t}
$$



## Relative Intensity After 24 hrs

The relative intensity of radiation after 24 hours

$$
\begin{aligned}
\gamma & =0.9998 \times e^{-0.1151(24)} \\
& =6.3160 \times 10^{-2}
\end{aligned}
$$

This result implies that only

$$
\frac{6.316 \times 10^{-2}}{0.9998}
$$

radioactive intensity is left after 24 hours.

## Linearization of data I

To find the constants of many nonlinear models, it results in solving simultaneous nonlinear equations. For mathematical convenience, some of the data for such models can be linearized. For example, the data for an exponential model can be linearized.

As shown in the previous example, many chemical and physical processes are governed by the equation,

$$
y=a e^{b x}
$$

Taking the natural $\log$ of both sides yields,

$$
\ln y=\ln a+b x
$$

Let $z=\ln y$ and $a_{0}=\ln a$
We now have a linear regression model where

$$
z=a_{0}+a_{1} x
$$

(implying) $\quad a=e^{a_{o}}$ with $a_{1}=b$

## Linearization of data II

Using linear model regression methods,

$$
\begin{aligned}
& a_{1}=\frac{n \sum_{i=1}^{n} x_{i} z_{i}-\sum_{i=1}^{n} x_{i} \sum_{i=1}^{n} z_{i}}{n \sum_{i=1}^{n} x_{i}^{2}-\left(\sum_{i=1}^{n} x_{i}\right)^{2}} \\
& a_{0}=\bar{z}-a_{1} \bar{x}
\end{aligned}
$$

Once $a_{o}, a_{1}$ are found, the original constants of the model are found as

$$
\begin{aligned}
& b=a_{1} \\
& a=e^{a_{0}}
\end{aligned}
$$

## Example - Linearization of data I

Same example of radioactive decay as before

| t (hrs) | 0 | 1 | 3 | 5 | 7 | 9 |
| :---: | :--- | :--- | :--- | :--- | :--- | :--- |
| $\gamma$ | 1.000 | 0.891 | 0.708 | 0.562 | 0.447 | 0.355 |

Exponential model given as,


$$
\begin{gathered}
\gamma=A e^{\lambda t} \\
\ln (\gamma)=\ln (A)+\lambda t
\end{gathered}
$$

Assuming $z=\ln \gamma, a_{o}=\ln (A)$ and $a_{1}=\lambda$ we obtain

$$
z=a_{0}+a_{1} t
$$

This is a linear relationship between
$z$ and $t$

## Example - Linearization of data II

Using this linear relationship, we can calculate $a_{0}, a_{1}$ where

$$
a_{1}=\frac{n \sum_{i=1}^{n} t_{i} z_{i}-\sum_{i=1}^{n} t_{i} \sum_{i=1}^{n} z_{i}}{n \sum_{i=1}^{n} t_{1}^{2}-\left(\sum_{i=1}^{n} t_{i}\right)^{2}} \quad \text { and } \quad \begin{aligned}
& a_{0}=\bar{z}-a_{1} \bar{t} \\
& \lambda
\end{aligned}=a_{1}, ~ A=e^{a_{0}} .
$$

Table. Summation data for linearization of data model

| $\boldsymbol{i}$ | $t_{i}$ | $\gamma_{i}$ | $z_{i}=\ln \gamma_{i}$ | $t_{i} z_{i}$ | $t_{i}^{2}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 0 | 1 | 0.00000 | 0.0000 | 0.0000 |
| 2 | 1 |  | 0.891 | -0.11541 | -0.11541 |
| 3 | 3 | 0.708 | -0.34531 | -1.0359 | 9.0000 |
| 4 | 5 | 0.562 | -0.57625 | -2.8813 | 25.000 |
| 5 | 7 | 0.447 | -0.80520 | -5.6364 | 49.000 |
| 6 | 9 | 0.355 | -1.0356 | -9.3207 | 81.000 |
| $\boldsymbol{\sum}$ | 25.000 |  | -2.8778 | -18.990 | 165.00 |

With $n=6$

$$
\begin{aligned}
& \sum_{i=1}^{6} t_{i}=25.000 \\
& \sum_{i=1}^{6} z_{i}=-2.8778 \\
& \sum_{i=1}^{6} t_{i} z_{i}=-18.990 \\
& \sum_{i=1}^{6} t_{i}^{2}=165.00
\end{aligned}
$$

## Example - Linearization of Data III

Calculating $a_{0}, a_{1}$

$$
\begin{aligned}
& a_{1}=\frac{6(-18.990)-(25)(-2.8778)}{6(165.00)-(25)^{2}}=-0.11505 \\
& a_{0}=\frac{-2.8778}{6}-(-0.11505) \frac{25}{6}=-2.6150 \times 10^{-4}
\end{aligned}
$$

Since

$$
\begin{aligned}
a_{0} & =\ln (A) \\
A & =e^{a_{0}} \\
& =e^{-2.6150 \times 10^{-4}}=0.99974
\end{aligned}
$$

Resulting model is $\gamma=0.99974 \times e^{-0.11505 t}$
also

$$
\lambda=a_{1}=-0.11505
$$



## Example - Linearization of Data IV

The regression formula is

$$
\gamma=0.99974 \times e^{-0.11505 t}
$$

The Half life of Technetium 99 is when $\quad \gamma=\left.\frac{1}{2} \gamma\right|_{t=0}$

$$
\begin{aligned}
& 0.99974 \times e^{-0.11505 t}=\frac{1}{2}(0.99974) e^{-0.11505(0)} \\
& e^{-0.11508 t}=0.5 \\
& -0.11505 t=\ln (0.5) \\
& t=6.0248 \text { hours }
\end{aligned}
$$

Literature value: 6.01 hours

## Comparison

Comparison of exponential model with and without data linearization:

|  | With data linearization | Without data linearization |
| :---: | :---: | :---: |
| A | 0.99974 | 0.99983 |
| $\lambda$ | -0.11505 | -0.11508 |
| Half-Life (hrs) | 6.0248 | 6.0232 |
| Relative intensity <br> after 24 hrs. | $6.3200 \times 10^{-2}$ | $6.3160 \times 10^{-2}$ |

The values are very similar so data linearization was suitable to find the constants of the nonlinear exponential model in this case.

## References

- This script is based on http://numericalmethods.eng.usf.edu by Autar Kaw, Jai Paul and Numerical Recipes (2nd/3rd Edition) by Press et al., Cambridge University Press http://www.nr.com/oldverswitcher.html


## Statistics and Fitting

Return by 9:15 a.m. tomorrow

## Assignment for the Afternoon / Homework

All data files required for this exercise sheet can be downloaded from the UKNum home-
page https://www2.mpia-hd.mpg.de/~klahr/UKNUM2022.html

- Exercise 1, 4 points: Statistics (I).

The data in the file dice.dat (available on the UKNum homepage) represent the results of experiments rolling two dice and adding the pips ("Augenzahl"). The first and second column in the file give the number of the experiment and the result, respectively. (a) Write a program code which reads in the data file and calculates the mean, the median and the standard deviation of the data. What is the relative difference between the mean and the median?

- Exercise 2, 6 points: Statistics (II).

The data in the file grades.dat represents the grades of an exam (with results between 0 (worst) and 100 (best). The number in the first column stands for the student's number, with his grade given in the second column. (a) Find the mean, the median and the standard deviation of the data. (b) What is the relative difference between the mean and the median? How does it compare to the result from 1? What does it tell you about the symmetry of the underlying parent distribution of the data?

- Exercise 3, 10 points: Least-Squares Fit to a Line

A student has measured the potential difference (in volts) along a conducting nickelsilver wire using an analog voltmeter. The voltage is measured between the negative end of the wire and various positions along the wire (in cm ). Uncertainties in the positions are less than 1 mm and can be neglected. Assume the the uncertainty in the voltage measurement to be the same for each measurement (i.e., set $\sigma_{i}=$ 1.0). (a) Write a program code to read in the data in the file data1.dat and fit a line $f(x)=a+b x$ through the data using a Least-Squares Fit as presented in the Lecture. (b) Estimate the standard deviation of an individual measurement using

$$
\sigma^{2} \sim s^{2}=\frac{1}{N-2} \sum\left(y_{i}-a-b x_{i}\right)^{2} .
$$

Continued on next page/backside

- Optional Task: Least-Squares Fit to an Exponential

The file data2.dat contains the measurements of an exponential decay, e.g., of a radioactive element (in some arbitrary units). The last column gives the (constant) uncertainty in each measurement. Fit an exponential function $f(x)=a \exp (b x)$ to the data using (a) the direct exponential fit as presented in the Lecture. Use a bracketing root-finding algorithm, e.g., bisection. In the current case, initial guesses of $b_{1}=-1$ and $b_{2}=0$ do bracket the root $b_{\star}$. (b) Fit a straight line to the linearized equation using a least-squares fit. Compare the result to the exponential fit in (a). Plot the results of (a) and (b) on linear and on logarithmic scale ( $\log y$ vs. $x$ ) including the errorbars.

